## NOTE ON A LIMIT THEOREM

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A simple derivation of a limiting distribution of Logan, Mallows, Rice and Shepp is found. This is the limiting distribution of a sequence of partial sums of independent, identically distributed random variables in the domain of attraction of a stable law, normalized by the partial sums of the absolute values of these variables.

Let  $X_1, X_2, \cdots$  be independent, identically distributed random variables, belonging to the domain of attraction of a stable distribution with index  $\alpha$ ,  $0 < \alpha < 1$ . In [2], Logan, Mallows, Rice and Shepp found, among many other interesting results, the limiting distribution of

(1) 
$$S_n = (\sum_{i=1}^n X_i) / (\sum_{i=1}^n |X_i|).$$

They did this by calculating the joint limiting characteristic function of the numerator and denominator in (1). In this note we utilize a property of stable laws, perhaps of independent interest, to obtain a simple proof of their result.

Write  $X_i = X_i^+ - X_i^-$ , where  $X_i^+$  and  $X_i^-$  are positive random variables:  $X_i^+ = \max(X_i, 0), X_i^- = -\min(X_i, 0).$ 

LEMMA. Let  $X_i$  belong to the domain of attraction of a stable law of index  $\alpha$ ,  $0 < \alpha < 1$ . Then there exist positive, independent stable laws,  $U^+$ ,  $U^-$  of index  $\alpha$ , and normalizing constants  $b_n$ , such that the joint limiting distribution of

$$(2) U_n^+ = b_n \sum_{i=1}^n X_i^+$$

$$(3) U_n^- = b_n \sum_{i=1}^n X_i^-$$

is that of  $(U^+, U^-)$ .

In particular  $X_i^+$  and  $X_i^-$  are attracted to the domain of a common stable law U.

To prove the lemma, by Doeblin's condition (cf. [1], page 175) we have

$$(4) P(X_i^+ > x) + P(X_i^- > x) \sim x^{-\alpha} L(x), x \to \infty$$

(5) 
$$\frac{P(X_i^- > x)}{P(X_i^+ > x)} \to \frac{C_1}{C_2}, \qquad x \to \infty$$

where L(x) is a slowly varying function, and  $C_1 \ge 0$ ,  $C_2 \ge 0$ ,  $C_1 + C_2 > 0$ . If  $C_2 = 0$ , then (5) is to be interpreted in the obvious way.

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The two "equations" (4) and (5) can be solved to give

$$P(X_i^- > x) \sim \frac{C_1}{C_1 + C_2} x^{-\alpha} L(x) , \qquad x \to \infty$$

$$P(X_i^+ > x) \sim \frac{C_2}{C_1 + C_2} x^{-\alpha} L(x),$$
  $x \to \infty$ 

and in addition, if  $\xi_1 > 0$ ,  $\xi_2 > 0$ , the random variables  $Y_i = \xi_1 X_i^- + \xi_2 X_i^+$  belong to the domain of attraction of U also, with the same normalizing constants  $b_n$ . This follows from the fact that for  $Z_i$  in the domain of attraction of a stable law the scale constants  $b_n$  can be chosen to satisfy  $P(Z_i > 1/b_n) \sim c/n$ , and from

$$\begin{split} P(Y_i > x) &= P(X_i^- > x/\xi_1) + P(X_i^+ > x/\xi_2) \\ &\sim \left(\frac{C_1 \xi_1^{\alpha}}{C_1 + C_2} + \frac{C_2 \xi_2^{\alpha}}{C_1 + C_2}\right) x^{-\alpha} L(x) \; . \end{split}$$

It now follows that the joint characteristic function of  $U_n^+$  and  $U_n^-$ , namely  $E(\exp(ib_n\sum_{i=1}^n Y_i))$ , converges to the product of the limiting characteristic functions of  $U_n^+$  and  $U_n^-$ . It follows also that the same conclusion is reached for arbitrary real  $\xi_1$ ,  $\xi_2$ , and the lemma is proved.

Since from (1)

$$S_n = \frac{U_n^+ - U_n^-}{U_n^+ + U_n^-}$$

where  $U_n^+$  and  $U_n^-$  are given by (2) and (3),  $S_n$  converges in distribution to

$$S = \frac{U^+ - U^-}{U^+ + U^-}$$
.

To find the distribution of S we shall use the following result. Let X be a random variable with a characteristic function  $\Psi(\xi) = E(\exp(i\xi X))$  and such that P(X=0)=0,  $E(|X|)<\infty$ . Then

(6) 
$$P(X > 0) = \frac{1}{2} + \lim_{T \to \infty} \frac{1}{\pi} \int_0^T \operatorname{Im} \Psi(\xi) \frac{d\xi}{\xi}.$$

The same formula holds if  $E(|X|) = \infty$ , but then the lower limit in (6) is also singular. (6) can be proved by using the well-known formula

$$\lim_{T\to\infty}\frac{2}{\pi}\,\int_0^T\,\frac{\sin\,at}{t}\,dt=\,\mathrm{sgn}\,a\,.$$

Let  $\varphi_+(z)$  and  $\varphi_-(z)$  be the characteristic functions of the two independent random variables  $U^+$  and  $U^-$ , so that

$$\varphi_{\pm}(z) = \exp\left\{-c_{\pm}(\cos\frac{1}{2}\pi\alpha + i\operatorname{sgn}z\sin\frac{1}{2}\pi\alpha)|z|^{\alpha}\right\}$$

where 
$$c_{-} = C_{1}/(C_{1} + C_{2})$$
,  $c_{+} = C_{2}/(C_{1} + C_{2})$ .

Then the distribution function of S is

$$F(x) = P(S < x) = P\left(\frac{U^{+} - U^{-}}{U^{+} + U^{-}} < x\right)$$

$$= P(U^{-}(1+x) - U^{+}(1-x) > 0)$$

$$= \frac{1}{2} + \frac{1}{\pi} \int_{0}^{\infty} \text{Im} \left[\varphi_{-}((1+x)z)\varphi_{+}(-(1-x)z)\right] \frac{dz}{z}$$

$$= \frac{1}{2} + \frac{1}{\pi} \int_{0}^{\infty} e^{-az^{\alpha}} \sin bz^{\alpha} \frac{dz}{z},$$

where, as easy algebra shows,

$$a = \cos \frac{1}{2}\pi\alpha (c_{-}(1+x)^{\alpha} + c_{+}(1-x)^{\alpha}),$$
  

$$b = -\sin \frac{1}{2}\pi\alpha (c_{-}(1+x)^{\alpha} - c_{+}(1-x)^{\alpha}).$$

The preceding integral is easy to evaluate, giving

$$F(x) = \frac{1}{2} - \frac{1}{\pi \alpha} \tan^{-1} \frac{b}{a}$$

$$= \frac{1}{2} + \frac{1}{\pi \alpha} \tan^{-1} \left( \left[ \frac{c_{-}(1+x)^{\alpha} - c_{+}(1-x)^{\alpha}}{c_{-}(1+x)^{\alpha} + c_{+}(1-x)^{\alpha}} \right] \tan \frac{1}{2} \pi \alpha \right), \qquad |x| < 1.$$

For  $\alpha = \frac{1}{2}$ ,  $c_+ = c_-$ , this becomes  $F(x) = \frac{1}{2} + 1/\pi \sin^{-1} x$ , a form of the "arcsin law."

## REFERENCES

- [1] GNEDENKO, B. V. and KOLMOGOROV, A. N. (1954). Limit Distributions for Sums of Independent Random Variables. Addison-Wesley, Cambridge.
- [2] LOGAN, B. F., MALLOWS, C. L., RICE, S. O., and SHEPP, L. A. (1973). Limit distributions of self-normalized sums. *Ann. Probability* 1 788-809.

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