

ON THE NON-CENTRAL DISTRIBUTIONS OF TWO TEST CRITERIA IN MULTIVARIATE ANALYSIS OF VARIANCE

BY C. G. KHATRI AND K. C. S. PILLAI¹

Gujarat University and Purdue University

1. Introduction and summary. Let $\mathbf{X}$ be a $p \times f_2$ matrix variate ($p \leq f_2$) and $\mathbf{Y}$ a $p \times f_1$ matrix variate ($p \leq f_1$) and the columns be all independently normally distributed with covariance matrix Σ , $E(\mathbf{X}) = \mathbf{M}$ and $E(\mathbf{Y}) = \mathbf{0}$. Let $0 < l_1 \leq \dots \leq l_p < 1$ be the ordered characteristic roots of

$$(1.1) \quad |\mathbf{X}\mathbf{X}' - l(\mathbf{Y}\mathbf{Y}' + \mathbf{X}\mathbf{X}')| = 0$$

and $\omega_1, \dots, \omega_p$ those of

$$|\mathbf{M}\mathbf{M}' - \omega\Sigma| = 0,$$

then the joint density function of $l_1, \dots, l_p$ is given by Constantine [1], James [2] in the form

$$(1.2) \quad \exp(-\frac{1}{2} \text{tr } \Omega) {}_1F_1(\frac{1}{2}\nu; \frac{1}{2}f_2; \frac{1}{2}\Omega, \mathbf{L}) f_0(l_1, \dots, l_p),$$

where

$$(1.3) \quad f_0(l_1, \dots, l_p) = C(p, f_1, f_2) \prod_{i=1}^p \{l_i^{\frac{1}{2}(f_2-p-1)} (1-l_i)^{\frac{1}{2}(f_1-p-1)}\} \alpha_p(\mathbf{L}),$$

$$\mathbf{L} = \mathbf{X}'(\mathbf{Y}\mathbf{Y}' + \mathbf{X}\mathbf{X}')^{-1}\mathbf{X}, \quad \Omega = \mathbf{M}'\Sigma^{-1}\mathbf{M}, \quad \nu = f_1 + f_2,$$

${}_1F_1$ is the hypergeometric function of matrix argument defined in [2] given by

$$\sum_{k=0}^{\infty} \sum_{\kappa} (\frac{1}{2}\nu)_{\kappa} C_{\kappa}(\frac{1}{2}\Omega) C_{\kappa}(\mathbf{L}) / (\frac{1}{2}f_2)_{\kappa} C_{\kappa}(\mathbf{I}_p) k!$$

and where

$$C(p, f_1, f_2) = \pi^{\frac{1}{2}p^2} \Gamma_p(\frac{1}{2}\nu) / \{\Gamma_p(\frac{1}{2}f_1) \Gamma_p(\frac{1}{2}f_2) \Gamma_p(\frac{1}{2}p)\}$$

and

$$\alpha_p(\mathbf{L}) = \prod_{i>j} (l_i - l_j).$$

In this paper the distribution of Pillai's $V^{(p)}$ criterion which is the trace of $\mathbf{L}$, [5], [6] and that of Roy's largest root criterion, l_p , [8], [10], have been obtained in series forms and certain constants involved in the series tabulated. In addition, the first four moments of $V^{(p)}$ are also obtained in the linear case, illustrating further use of some of the tabulations.

2. Distribution of $V^{(p)}$. First a lemma is proved which will be used later in the section.

Received 20 March 1967.

¹ The work of this author was supported by the National Science Foundation, Grant No. GP-4600.

LEMMA 1. If A_κ denotes the integral

$$(2.1) \quad \int_{\mathfrak{D}} |\mathbf{Z}|^{\frac{1}{2}(f_2-p-1)} \alpha_p(\mathbf{Z}) C_\kappa(\mathbf{Z}) dz_1, \dots, dz_{p-1},$$

where

$$\mathbf{Z} = \text{diag}(z_1, \dots, z_p), \quad z_p = 1 - z_1 - \dots - z_{p-1}$$

and $\mathfrak{D}$ is given by

$$(2.2) \quad \mathfrak{D}: \{0 < z_1 < z_2 < \dots < z_{p-2} < z_{p-1} < z_p = 1 - z_1 - \dots - z_{p-2} - z_{p-1}\},$$

then

$$(2.3) \quad A_\kappa = \{(\frac{1}{2}f_2)_\kappa \Gamma_p(\frac{1}{2}f_2) \Gamma_p(\frac{1}{2}p) C_\kappa(\mathbf{I}_p)\} / \{\pi^{\frac{1}{2}p^2} (\frac{1}{2}f_2p)_\kappa \Gamma(\frac{1}{2}f_2p)\}.$$

PROOF. Consider the null distribution of $T = \sum_{i=1}^p c_i$, $c_i = l_i/(1 - l_i)$, $i = 1, \dots, p$, given by Constantine [1] in the form

$$(2.4) \quad \Gamma_p(\frac{1}{2}\nu) \{\Gamma(\frac{1}{2}f_2p) \Gamma_p(\frac{1}{2}f_1)\}^{-1} T^{\frac{1}{2}f_2p-1} \sum_{k=0}^{\infty} (-T)^k \{k! (\frac{1}{2}f_2p)_k\}^{-1} \cdot \sum_{\kappa} (\frac{1}{2}\nu)_\kappa (\frac{1}{2}f_2)_\kappa C_\kappa(\mathbf{I}_p),$$

for any $|T| < 1$. Now, from $f_0(l_1, \dots, l_p)$ in (1.3), the joint density of $c_1, \dots, c_p$ can be obtained as

$$(2.5) \quad C(p, f_1, f_2) \prod_{i=1}^p \{c_i^{\frac{1}{2}(f_2-p-1)} (1 + c_i)^{-\frac{1}{2}\nu}\} \alpha_p(\mathbf{C}), \quad 0 < c_1 \leq \dots \leq c_p < \infty,$$

where $\mathbf{C} = \text{diag}(c_1, \dots, c_p)$. Expanding

$$(2.6) \quad \prod_{i=1}^p (1 + c_i)^{-\frac{1}{2}\nu} = \sum_{k=0}^{\infty} (-1)^k \sum_{\kappa} (\frac{1}{2}\nu)_\kappa C_\kappa(\mathbf{C})/k!,$$

transforming $z_i = c_i/T$ and integrating z_i 's in the region $\mathfrak{D}$ given in (2.2) we get

$$(2.7) \quad C(p, f_1, f_2) T^{\frac{1}{2}f_2p-1} \sum_{k=0}^{\infty} (-T)^k (k!)^{-1} \sum_{\kappa} (\frac{1}{2}\nu)_\kappa A_\kappa,$$

where A_κ is given in (2.1). Hence equating the coefficients of $(\frac{1}{2}\nu)_\kappa$ in (2.4) and (2.7) we get (2.3). Hence the lemma.

Now consider the distribution of $l_1, \dots, l_p$ in (1.2). First note that

$$(2.8) \quad |\mathbf{I} - \mathbf{L}|^{\frac{1}{2}(f_1-p-1)} C_\kappa(\mathbf{L}) = \sum_{n=0}^{\infty} \sum_{\eta} (\frac{1}{2}(p+1-f_1))_{\eta} C_{\eta}(\mathbf{L}) \mathbf{C}_{\kappa}(\mathbf{L})/n! \\ = \sum_{n=0}^{\infty} \sum_{\eta} \sum_{\delta} (\frac{1}{2}(p+1-f_1))_{\eta} g_{\kappa, \eta}^{\delta} C_{\delta}(\mathbf{L})/n!$$

where $g_{\kappa, \eta}^{\delta}$ is the coefficient of $C_{\delta}(\mathbf{L})$ in the product $C_{\kappa}(\mathbf{L}) C_{\eta}(\mathbf{L})$, $\delta = (\delta_1, \dots, \delta_p)$, $\delta_1 \geq \dots \geq \delta_p \geq 0$ and $\sum_{i=1}^p \delta_i = n + k = d$. Using (2.8) in (1.2) and integrating $l_1, \dots, l_p$ over the surface $\sum_{i=1}^p l_i = V^{(p)}$, we get the density of $V^{(p)}$ as

$$(2.9) \quad C(p, f_1, f_2) \exp(-\frac{1}{2} \text{tr } \mathbf{\Omega}) \sum_{k=0}^{\infty} \sum_{\kappa} \sum_{\eta=0}^{\infty} \sum_{\eta} \sum_{\delta} (\frac{1}{2}\nu)_{\kappa} (\frac{1}{2}(p+1-f_1))_{\eta} \\ \cdot C_{\kappa}(\frac{1}{2}\mathbf{\Omega}) g_{\kappa, \eta}^{\delta} A_{\delta} \{V^{(p)}\}^{\frac{1}{2}f_2p+d-1} / \{(\frac{1}{2}f_2)_{\kappa} C_{\kappa}(\mathbf{I}_p) k! n!\},$$

where A_{δ} is defined by (2.1) whose value is obtained from (2.3) by putting $\kappa = \delta$. The series (2.9) is convergent for $0 < V^{(p)} < 1$. Tabulations of the g -coefficients are presented in Table 1 for various values of the three arguments.

3. The moments of $V^{(p)}$ in the linear case. Consider now the evaluation of the moments of $V^{(p)}$ in the linear case, i.e. when there is only one non-zero population root, say, ω . This will also serve to illustrate further uses of Table 1. For this, as the authors have shown in a previous paper, [3], we may write

$$(3.1) \quad V^{(p)} - 1 = \text{tr } \mathbf{L}_{22} - (1 - l_{11})[(1 - \mathbf{u}'\mathbf{u}) + \mathbf{u}'\mathbf{L}_{22}\mathbf{u}] \\ = \sum_{i=1}^{p-1} c_i' - (1 - l_{11})[(1 - \mathbf{u}'\mathbf{u}) + \sum_{i=1}^{p-1} c_i' u_i^2],$$

where l_{11} , $\mathbf{u}:(p-1 \times 1)$ and $\mathbf{L}_{22}:(p-1) \times (p-1)$ are independently distributed and whose distributions are given in [3], [4], and c_i' ($i = 1, \dots, p-1$) are the characteristic roots of $\mathbf{L}_{22}$. It may be pointed out that the density of l_{11} alone involves ω . Now let $l_{11,0}$ be a variate whose distribution is the same as that of l_{11} when $\omega = 0$ and $V_0^{(p)}$ be the corresponding $V^{(p)}$ statistic. We may note that

$$(3.2) \quad x_1 = E(1 - l_{11,0}) - E(1 - l_{11}) = f_1 \delta(\nu),$$

$$(3.3) \quad x_2 = E(1 - l_{11,0})^2 - E(1 - l_{11})^2 = \frac{1}{2} f_1 (f_1 + 2) \Delta_1,$$

$$(3.4) \quad x_3 = E(1 - l_{11,0})^3 - E(1 - l_{11})^3 = \frac{1}{8} f_1 (f_1 + 2)(f_1 + 4) \Delta_2,$$

and

$$(3.5) \quad x_4 = E(1 - l_{11,0})^4 - E(1 - l_{11})^4 = \frac{1}{48} f_1 (f_1 + 2)(f_1 + 4)(f_1 + 6) \Delta_3,$$

where

$$(3.6) \quad \delta(\nu) = [\omega/2\nu] \exp(-\frac{1}{2}\omega) \sum_{i=0}^{\infty} \{(\frac{1}{2}\omega)^i / [i! (\frac{1}{2}\nu + i + 1)]\},$$

$$(3.7) \quad \Delta_1 = \delta(\nu) - \delta(\nu + 2), \quad \Delta_2 = \delta(\nu) - 2\delta(\nu + 2) + \delta(\nu + 4)$$

$$\text{and} \quad \Delta_3 = \delta(\nu) - 3\delta(\nu + 2) + 3\delta(\nu + 4) - \delta(\nu + 6).$$

Now from (3.1) we can write

$$(3.8) \quad E(V^{(p)} - 1) = E(V_0^{(p)} - 1) + x_1 E(\beta),$$

$$(3.9) \quad E(V^{(p)} - 1)^2 = E(V_0^{(p)} - 1)^2 - x_2 E(\beta^2) + 2x_1 E(\beta\alpha),$$

$$(3.10) \quad E(V^{(p)} - 1)^3 = E(V_0^{(p)} - 1)^3 + \sum_{i=1}^3 \binom{3}{i} (-1)^{i-1} E(\beta^i \alpha^{3-i}) x_i,$$

$$(3.11) \quad E(V^{(p)} - 1)^4 = E(V_0^{(p)} - 1)^4 + \sum_{i=1}^4 \binom{4}{i} (-1)^{i-1} E(\beta^i \alpha^{4-i}) x_i$$

where

$$(3.12) \quad \alpha = \text{tr } \mathbf{L}_{22} \quad \text{and} \quad \beta = 1 - \mathbf{u}'\mathbf{u} + \mathbf{u}'\mathbf{L}_{22}\mathbf{u}.$$

Now, the expected values on the right side of (3.8)–(3.11) which are the coefficients of the x 's have been obtained in terms of functions of $\text{tr } \mathbf{L}_{22}$ i.e. the i th elementary symmetric function in the characteristic roots, $c_1', \dots, c_{p-1}'$,

TABLE 1
Values of $g_{\alpha,\beta}$

$C_{\alpha}(\cdot)$	$C_{\beta}(\cdot)$	$k = 2$		$k = 3$			$k = 4$					
		$C_{(2)}$	$C_{(1^2)}$	$C_{(3)}$	$C_{(2,1)}$	$C_{(1^3)}$	$C_{(4)}$	$C_{(3,1)}$	$C_{(2^2)}$	$C_{(2,1^2)}$	$C_{(1^4)}$	
$C_{(1)}$	$C_{(1)}$	1	1									
$C_{(1)}$	$C_{(2)}$			1	4/9							
$C_{(1)}$	$C_{(1^2)}$				5/9	1						
$C_{(1)}$	$C_{(3)}$						1	3/10	1	5/8		
$C_{(1)}$	$C_{(2,1)}$							7/10		3/8	1	
$C_{(1)}$	$C_{(1^3)}$											
$C_{(2)}$	$C_{(2)}$						1	2/9	4/9	5/18		
$C_{(2)}$	$C_{(3)}$							7/18	5/9	4/9		
$C_{(2)}$	$C_{(1^2)}$										1	
$C_{(2)}$	$C_{(1^3)}$											

$C_{\alpha}(\cdot)$	$C_{\beta}(\cdot)$	$k = 5$									
		$C_{(5)}$	$C_{(4,1)}$	$C_{(3,2)}$	$C_{(3,1^2)}$	$C_{(2^2,1)}$	$C_{(2,1^3)}$	$C_{(1^5)}$			
$C_{(1)}$	$C_{(4)}$	1	8/35	8/15	7/15	7/15	1/3	2/3	18/25	1	
$C_{(1)}$	$C_{(3,1)}$		27/35	7/15			8/15		7/25		
$C_{(1)}$	$C_{(2^2)}$										
$C_{(1)}$	$C_{(2,1^2)}$										
$C_{(1)}$	$C_{(1^4)}$										
$C_{(1)}$	$C_{(3)}$										
$C_{(1)}$	$C_{(2,1)}$										
$C_{(2)}$	$C_{(2)}$	1	4/25	4/25	7/30	1/3	7/30	1/5	1/5		
$C_{(2)}$	$C_{(3)}$		27/50	28/75	1/5		7/50				
$C_{(2)}$	$C_{(1^2)}$										
$C_{(2)}$	$C_{(1^3)}$										
$C_{(2)}$	$C_{(2,1)}$		3/10				32/75	5/12	9/20	7/20	1
$C_{(2)}$	$C_{(1^2)}$							1/4			
$C_{(2)}$	$C_{(1^3)}$										

		$k = 6$											
$C_{\alpha}(\cdot)$	$C_{\beta}(\cdot)$	$C_{(6)}$	$C_{(6,1)}$	$C_{(4,2)}$	$C_{(4,3)}$	$C_{(4^4)}$	$C_{(3,2,1)}$	$C_{(3,1^5)}$	$C_{(2^5)}$	$C_{(2^3,1^3)}$	$C_{(3,1^4)}$	$C_{(1^6)}$	$C_{(1^5)}$
$C_{(1)}$	$C_{(6)}$	1	$\frac{1}{54}$										
$C_{(1)}$	$C_{(4,1)}$		$\frac{22}{27}$										
$C_{(1)}$	$C_{(3,2)}$			$\frac{2}{5}$	$\frac{3}{8}$	1	$\frac{7}{27}$	$\frac{4}{7}$		1	$\frac{1}{2}$		
$C_{(1)}$	$C_{(3,1^2)}$			$\frac{3}{5}$	$\frac{5}{8}$		$\frac{10}{27}$				$\frac{1}{2}$		
$C_{(1)}$	$C_{(2^2,1)}$						$\frac{10}{27}$	$\frac{3}{7}$					
$C_{(1)}$	$C_{(2,1^3)}$												
$C_{(1)}$	$C_{(1^6)}$												
$C_{(1)}$	$C_{(4)}$	1	$\frac{8}{63}$	$\frac{16}{175}$									1
$C_{(2)}$	$C_{(3,1)}$		$\frac{22}{35}$	$\frac{8}{35}$	$\frac{1}{5}$	$\frac{8}{15}$	$\frac{56}{405}$			$\frac{1}{3}$			
$C_{(2)}$	$C_{(2^2)}$			$\frac{7}{25}$			$\frac{7}{81}$						
$C_{(2)}$	$C_{(2,1^2)}$				$\frac{1}{3}$		$\frac{16}{81}$	$\frac{16}{75}$			$\frac{4}{15}$	$\frac{7}{45}$	
$C_{(2)}$	$C_{(1^4)}$							$\frac{3}{25}$					
$C_{(1^2)}$	$C_{(4)}$		$\frac{11}{45}$					$\frac{4}{15}$					
$C_{(1^2)}$	$C_{(3,1)}$			$\frac{2}{5}$	$\frac{3}{35}$		$\frac{14}{81}$						
$C_{(1^2)}$	$C_{(2^2)}$				$\frac{8}{21}$	$\frac{7}{15}$	$\frac{64}{405}$			$\frac{2}{3}$	$\frac{32}{75}$	$\frac{14}{25}$	
$C_{(1^2)}$	$C_{(2,1^2)}$						$\frac{20}{81}$	$\frac{2}{5}$			$\frac{7}{50}$	$\frac{64}{225}$	1
$C_{(1^2)}$	$C_{(1^4)}$												
$C_{(3)}$	$C_{(6)}$	1	$\frac{3}{25}$	$\frac{16}{25}$		$\frac{4}{25}$	$\frac{7}{75}$						
$C_{(3)}$	$C_{(2,1)}$		$\frac{11}{25}$	$\frac{27}{125}$	$\frac{27}{200}$								
$C_{(3)}$	$C_{(1^3)}$				$\frac{1}{8}$	$\frac{21}{25}$	$\frac{12}{25}$	$\frac{2}{25}$		$\frac{3}{4}$	$\frac{3}{8}$	$\frac{7}{20}$	
$C_{(2,1)}$	$C_{(2,1)}$			$\frac{63}{125}$	$\frac{12}{25}$		$\frac{1}{6}$	$\frac{3}{10}$		$\frac{9}{40}$	$\frac{7}{40}$	$\frac{3}{10}$	1
$C_{(2,1)}$	$C_{(1^3)}$							$\frac{27}{100}$		$\frac{1}{4}$			
$C_{(1^3)}$	$C_{(1^3)}$												

TABLE 1 (Cont'd)

$C_{\alpha}(\cdot)$		$C_{\beta}(\cdot)$	$k = 7$																	
$C(\eta)$	$C(\epsilon_1)$	$C(\epsilon_2)$	$C(\epsilon_1\epsilon_2)$	$C(\epsilon_1\epsilon_3)$	$C(\epsilon_1\epsilon_4)$	$C(\epsilon_1\epsilon_5)$	$C(\epsilon_1\epsilon_6)$	$C(\epsilon_1\epsilon_7)$	$C(\epsilon_2\epsilon_3)$	$C(\epsilon_2\epsilon_4)$	$C(\epsilon_2\epsilon_5)$	$C(\epsilon_2\epsilon_6)$	$C(\epsilon_2\epsilon_7)$	$C(\epsilon_3\epsilon_4)$	$C(\epsilon_3\epsilon_5)$	$C(\epsilon_3\epsilon_6)$	$C(\epsilon_3\epsilon_7)$	$C(\epsilon_4\epsilon_5)$	$C(\epsilon_4\epsilon_6)$	$C(\epsilon_4\epsilon_7)$
1	12/77 65/77	16/49 33/49	11/35 24/35	4/7 3/7	10/21 11/21	1/5 4/5	5/7 2/7	20/49 2/7	2/7	15/49	5/14	2/7 5/7	3/5 2/5	40/49 9/49	1					
1	20/189 130/189	80/1323 176/945 99/245	11/63 3/7	8/35 12/35	3/35 4/45 25/189 5/27	10/49 11/49	4/27 8/27	5/27 40/189	160/1323 20/189 6/49	2/7	4/21 5/63	2/9	8/63							
1	13/63	22/63	11/189 64/189	3/7	5/28 77/196	7/27	50/189 64/189	20/189 256/1323 25/126 15/98	18/49	5/14 5/14	3/10 7/18 4/45	40/63 5/21	1							

of $\mathbf{L}_{22}$ [4]. Thus (3.10) reduces to the form [4]

$$\begin{aligned}
 E(V^{(p)} - 1)^3 &= E(V_0^{(p)} - 1)^3 + \frac{1}{8}\Delta_2 f^{(3)} \\
 &+ \frac{3}{8}\{-3\delta(\nu) + 2\delta(\nu + 2) + \delta(\nu + 4)\}f^{(2)}E(\text{tr } \mathbf{L}_{22}) \\
 &+ \{3\delta(\nu) + 2\delta(\nu + 2) + 3\delta(\nu + 4)\}f^{(1)}E(\text{tr } \mathbf{L}_{22})^2 \\
 &+ \{\delta(\nu) + 2\delta(\nu + 2) + 5\delta(\nu + 4)\}E(\text{tr } \mathbf{L}_{22})^3 \\
 &+ 4\{\delta(\nu) + 2\delta(\nu + 2) - 3\delta(\nu + 4)\}E(\text{tr } \mathbf{L}_{22} \text{ tr}_2 \mathbf{L}_{22}) \\
 &+ 4\Delta_2\{-f^{(1)}E(\text{tr}_2 \mathbf{L}_{22}) + 2E(\text{tr}_3 \mathbf{L}_{22})\},
 \end{aligned}
 \tag{3.13}$$

where

$$\begin{aligned}
 f^{(i)} &= \prod_{j=1}^i (f_1 - p - 1 + 2j), \\
 E(\text{tr}_i \mathbf{L}_{22}) &= \binom{p-1}{i} \prod_{j=1}^i [(f_2 - j)/(\nu - j)] \\
 &\quad (i = 1, 2, \dots, p-1; \text{tr}_1 \mathbf{L}_{22} = \text{tr } \mathbf{L}_{22}),
 \end{aligned}
 \tag{3.14}$$

$$\begin{aligned}
 E(\text{tr } \mathbf{L}_{22} \text{ tr}_i \mathbf{L}_{22}) &= [E(\text{tr}_i \mathbf{L}_{22})/(\nu + 1)(\nu - i - 1)] \\
 &\cdot [(f_1 - p)\{f_2 - 1\}(p - 1) + 2i] \\
 &+ (p - 1)(f_2 - p - 1)(f_2 + 2p - i + 1) \\
 &+ 2p^3 - ip^2 + (i - 2)p + 2i]
 \end{aligned}
 \tag{3.15}$$

and

$$\begin{aligned}
 E(\text{tr } \mathbf{L}_{22})^3 &= E(\text{tr } \mathbf{L}_{22})[2^3(f_1 - f_2 + 1)f_1(\nu - 2p + 1)(\nu - p) \\
 &\cdot \{(\nu - 3)(\nu - 2)(\nu - 1)^2(\nu + 1)(\nu + 3)\}^{-1} + 3E(\text{tr } \mathbf{L}_{22})^2 \\
 &- 2\{E(\text{tr } \mathbf{L}_{22})\}^2].
 \end{aligned}
 \tag{3.16}$$

$$\begin{aligned}
 E(V^{(p)} - 1)^4 &= E(V_0^{(p)} - 1)^4 - \frac{1}{48}\Delta_3 f^{(4)} \\
 &+ \frac{1}{12}\{5\delta(\nu) - 9\delta(\nu + 2) + 3\delta(\nu + 4) + \delta(\nu + 6)\}f^{(3)} \\
 &\cdot E(\text{tr } \mathbf{L}_{22}) \\
 &+ \frac{3}{8}\{-5\delta(\nu) + 3\delta(\nu + 2) + \delta(\nu + 4) + \delta(\nu + 6)\}f^{(2)} \\
 &\cdot E(\text{tr } \mathbf{L}_{22})^2 \\
 &+ \frac{1}{4}\{5\delta(\nu) + 3\delta(\nu + 2) + 3\delta(\nu + 4) + 5\delta(\nu + 6)\}f^{(1)} \\
 &\cdot E(\text{tr } \mathbf{L}_{22})^3 \\
 &+ \frac{1}{12}\{5\delta(\nu) + 9\delta(\nu + 2) + 15\delta(\nu + 4) + 35\delta(\nu + 6)\} \\
 &\cdot E(\text{tr } \mathbf{L}_{22})^4 \\
 &+ \frac{3}{2}\{\delta(\nu) + \delta(\nu + 2) + 3\delta(\nu + 4) - 5\delta(\nu + 6)\} \\
 &\cdot E[(\text{tr } \mathbf{L}_{22})^2 \text{tr}_2 \mathbf{L}_{22}] \\
 &+ 3\Delta\{-f^{(1)}E(\text{tr } \mathbf{L}_{22} \text{ tr}_2 \mathbf{L}_{22}) + 2E(\text{tr } \mathbf{L}_{22} \text{ tr}_3 \mathbf{L}_{22})\} \\
 &+ \Delta_3\{f^{(2)}/2E(\text{tr}_2 \mathbf{L}_{22}) - 2f^{(1)}E(\text{tr}_3 \mathbf{L}_{22}) + 4E(\text{tr}_4 \mathbf{L}_{22}) \\
 &- 3E(\text{tr}_2 \mathbf{L}_{22})^2\},
 \end{aligned}
 \tag{3.17}$$

where

$$\Delta = \delta(\nu) - \delta(\nu + 2) - \delta(\nu + 4) + \delta(\nu + 6).$$

Now using Table 1,

$$(3.18) \quad E[(\text{tr } \mathbf{L}_{22})^2 \text{tr}_2 \mathbf{L}_{22}] = \frac{1}{2^4} E[7C_{(31)}(\mathbf{L}_{22}) + 10C_{(2^2)}(\mathbf{L}_{22}) \\ + 13C_{(21^2)}(\mathbf{L}_{22}) + 18C_{(1^4)}(\mathbf{L}_{22})].$$

Now using a theorem of Constantine (See Theorem 3 of [1]) the right side of (3.18) equals

$$(3.19) \quad \frac{1}{2^4} \Gamma_{p-1}(\frac{1}{2}f_1)[7b_1 + 10b_2 + 13b_3 + 18b_4],$$

where

$$\begin{aligned} b_1 &= \Gamma_{p-1}(t, (31))C_{(31)}(\mathbf{I})/\Gamma_{p-1}(v, (31)), \\ b_2 &= \Gamma_{p-1}(t, (2^2))C_{(2^2)}(\mathbf{I})/\Gamma_{p-1}(v, (2^2)), \\ b_3 &= \Gamma_{p-1}(t, (21^2))C_{(21^2)}(\mathbf{I})/\Gamma_{p-1}(v, (21^2)), \\ b_4 &= \Gamma_{p-1}(t, (1^4))C_{(1^4)}(\mathbf{I})/\Gamma_{p-1}(v, (1^4)), \end{aligned}$$

where

$$t = \frac{1}{2}(f_2 - 1) \quad \text{and} \quad v = \frac{1}{2}(\nu - 1).$$

Similarly from Table 1,

$$(3.20) \quad E(\text{tr}_2 \mathbf{L}_{22})^2 = \frac{1}{1^6} \Gamma_{p-1}(\frac{1}{2}f_1)[5b_2 + 4b_3 + 9b_4]$$

and

$$(3.21) \quad E(\text{tr}_1 \mathbf{L}_{22})^4 = \Gamma_{p-1}(\frac{1}{2}f_1) \sum_{i=0}^4 b_i,$$

where

$$b_0 = \Gamma_{p-1}(t, (4))C_{(4)}(\mathbf{I})/\Gamma_{p-1}(v, (4)).$$

It may be pointed out that alternate expressions for $E[(\text{tr } \mathbf{L}_{22})^2 \text{tr}_2 \mathbf{L}_{22}]$, $E(\text{tr}_2 \mathbf{L}_{22})$ and $E(\text{tr}_1 \mathbf{L}_{22})^4$ are available in [4]. The latter two can further be obtained from [7], [8], [9]. In addition, the first four moments of $V_0^{(p)}$ are available in Pillai [5], [7], [8] and the first two moments of $V^{(p)}$ in Khatri and Pillai [3] and hence (3.8) and (3.9) are not treated here separately.

4. Distribution of the largest root, l_p . In the density function of $l_1, \dots, l_p$ in (1.2) let us make the following transformation: $l_i/l_p = g_i, i = 1, 2, \dots, p-1$, $l_p = l_p$. Then the joint distribution of $(g_1, \dots, g_{p-1})$ and l_p is given by

$$(4.1) \quad C(p, f_1, f_2) \exp(-\frac{1}{2} \text{tr } \mathbf{\Omega}) l_p^{\frac{1}{2}pf_2-1} (1 - l_p)^{\frac{1}{2}(f_1-p-1)} \alpha_{p-1}(\mathbf{G}) |\mathbf{G}|^{\frac{1}{2}(f_2-p-1)} \\ \cdot |\mathbf{I}_{p-1} - l_p \mathbf{G}|^{\frac{1}{2}(f_1-p-1)} |\mathbf{I}_{p-1} - \mathbf{G}| \sum_{k=0}^{\infty} \sum_{\kappa} (\frac{1}{2}\nu)_{\kappa} C_{\kappa}(\frac{1}{2}\mathbf{\Omega}) C_{\kappa}(\mathbf{G}_1) l_p^{\kappa} / \\ \{(\frac{1}{2}f_2)_{\kappa} C_{\kappa}(\mathbf{I}_p) k!\},$$

TABLE 2
Values of $b_{\kappa, \eta}$

n	η	k = 1	k = 2		k = 3		
		κ					
		[1]	[2]	[1²]	[3]	[2 1]	[1³]
0		1	1		1		
1	[1]	1	2/3	4/3	3/5	12/5	
2	[2]		1		3/5	12/5	
3	[1²]			1		3/2	3/2
	[3]				1		
	[2 1]					1	
	[1³]						1

n	η	k = 4				
		κ				
		[4]	[3 1]	[2²]	[2 1²]	[1⁴]
0		1				
1	[1]	4/7	24/7			
2	[2]	18/35	16/7	16/5		
3	[1²]		2		4	
	[3]	4/7	24/7			
	[2 1]		8/9	8/9	20/9	
4	[1³]				12/5	8/5
	[4]	1				
	[3 1]		1			
	[2²]			1		
	[2 1²]				1	
	[1⁴]					1

n	η	k = 5						
		κ						
		[5]	[4 1]	[3 2]	[3 1²]	[2² 1]	[2 1³]	[1⁵]
0		1						
1	[1]	5/9	40/9					
2	[2]	10/21	8/3	48/7				
3	[1²]		5/2		15/2			
	[3]	10/21	8/3	48/7				
	[2 1]		1	16/9	25/9	40/9		
4	[1³]				30/7		40/7	
	[4]	5/9	40/9					
	[3 1]		3/4	4/3	35/12			
	[2²]			5/3		10/3		
	[2 1²]				25/21	5/3	15/7	
5	[1⁴]						10/3	
	[5]	1						
	[4 1]		1					
	[3 2]			1				
	[3 1²]				1			
	[2² 1]					1		
	[2 1³]					1		
	[1⁵]						1	

TABLE 2 (Cont'd.)

n	η	$k = 6$										
		κ										
		[6]	[5 1]	[4 2]	[4 1 ²]	[3 ²]	[3 2 1]	[3 1 ³]	[2 ³]	[2 ² 1 ²]	[2 1 ⁴]	[1 ⁶]
0		1										
1	[1]	6/11	60/11									
2	[2]	5/11	240/77	80/7								
	[1 ²]		3	12								
3	[3]	100/231	216/77	160/21		64/7						
	[2 1]		8/7	20/7	4		12					
	[1 ³]			20/3				40/3				
4	[4]	5/11	240/77	80/7								
	[3 1]		27/35	10/7	14/5	8/5	42/5					
	[2 ²]			5/2			15/2		5			
	[2 1 ²]				5/3		30/7	10/3		40/7		
	[1 ⁴]							15/2			15/2	
5	[5]	6/11	60/11									
	[4 1]		24/35	12/7	18/5							
	[3 2]			1		4/5	21/5					
	[3 1 ²]				14/15		12/5	8/3				
	[2 ² 1]						15/7		1	20/7		
	[2 1 ³]							3/2		12/5	21/10	
	[1 ⁵]										30/7	12/7
6	[6]	1										
	[5 1]		1									
	[4 2]			1								
	[4 1 ²]				1							
	[3 ²]					1						
	[3 2 1]						1					
	[3 1 ³]							1				
	[2 ³]								1			
	[2 ² 1 ²]									1		
	[2 1 ⁴]										1	
	[1 ⁶]											1

where $\mathbf{G} = \text{diag}(g_1, \dots, g_{p-1})$ and $\mathbf{G}_1 = \text{diag}(\mathbf{G}, 1)$. Now the distribution of l_p is obtained from (4.1) by integrating $g_1, \dots, g_{p-1}$ i.e. by evaluating the integral

$$(4.2) \quad B_k = \int \int_{0 \leq g_1 \leq \dots \leq g_{p-1} \leq 1} |\mathbf{G}|^{\frac{1}{2}(f_2 - p - 1)} \cdot |\mathbf{I}_{p-1} - \mathbf{G}|^{\alpha_{p-1}(\mathbf{G})} |\mathbf{I}_{p-1} - l_p \mathbf{G}|^{\frac{1}{2}(f_1 - p - 1)} C_\kappa(\mathbf{G}_1) d\mathbf{G}.$$

Now let us note the following results:

$$(4.3) \quad C_\kappa(\mathbf{G}_1) = \sum_{n=0}^k \sum_{\eta} b_{\kappa, \eta} C_\eta(\mathbf{G}),$$

where $b_{\kappa, \eta}$'s are constants depending on κ and η ,

$$(4.4) \quad |\mathbf{I}_p - l_p \mathbf{G}|^{\frac{1}{2}(f_1 - p - 1)} = \sum_{d=0}^{\infty} \sum_r \left(\frac{1}{2}(p + 1 - f_1)\right) \tau l_p^d C_r(\mathbf{G}) / d!.$$

As in (2.8)

$$(4.5) \quad C_\eta(\mathbf{G})C_\tau(\mathbf{G}) = \sum_{\delta} g_{\eta,\tau}^{\delta} C_{\delta}(\mathbf{G}).$$

Using (4.3)–(4.5) in (4.2) we get

$$(4.6) \quad B_{\kappa} = \sum_{n=0} \sum_{\eta} \sum_{d=0} \sum_{\tau} \sum_{\delta} b_{\kappa,\eta} g_{\eta,\tau}^{\delta} (\tfrac{1}{2}(p+1-f_1))_{\tau} l_p^d D_{\delta} / d!,$$

where

$$(4.7) \quad \begin{aligned} D_{\delta} &= \int \int_{0 \leq \theta_1 \leq \dots \leq \theta_{p-1} \leq 1} |\mathbf{G}|^{\frac{1}{2}(f_2-p-1)} |\mathbf{I}_{p-1} - \mathbf{G}| \alpha_{p-1}(\mathbf{G}) C_{\delta}(\mathbf{G}) d\mathbf{G} \\ &= \Gamma_{p-1}(\tfrac{1}{2}(p-1)) \Gamma_{p-1}(\tfrac{1}{2}(p+1)) \Gamma_{p-1}(\tfrac{1}{2}(f_2-1), \delta) C_{\delta}(\mathbf{I}_{p-1}) \\ &\quad \cdot [\pi^{\frac{1}{2}(p-1)^2} \Gamma_{p-1}(\tfrac{1}{2}(f_2+p), \delta)]^{-1}. \end{aligned}$$

Using (4.6) in (4.1) we get the distribution of l_p . Table 2 gives the values of $b_{\kappa,\eta}$ for various values of κ and η . If a_i denotes the i th elementary symmetric function of the roots of $\mathbf{G}_1$ i.e. $a_i = \text{tr}_i \mathbf{G}_1$, $i = 1, 2, \dots$, the $b_{\kappa,\eta}$ coefficients were obtained by using the zonal polynomials tabulated by James [2] in terms of the a_i 's and using the relation $\text{tr}_i \mathbf{G}_1 = \text{tr}_i \mathbf{G} + \text{tr}_{i-1} \mathbf{G}$.

The authors wish to thank Alan T. James and Anne Parkhurst for providing them with the tabulations of zonal polynomials of degrees seven to eleven. They also wish to express their thanks to Mrs. Louise Mao Lui, Statistics Section of Computer Sciences, Purdue University, for the excellent programming of the material for the computations in this paper carried out on IBM 7094, Purdue University's Computer Science's Center.

REFERENCES

- [1] CONSTANTINE, A. G. (1963). Some non-central distribution problems in multivariate analysis. *Ann. Math. Statist.* **34** 1270–1285.
- [2] JAMES, A. T. (1964). Distributions of matrix variates and latent roots derived from normal samples. *Ann. Math. Statist.* **35** 475–501.
- [3] KHATRI, C. G. and PILLAI, K. C. S. (1965). Some results on the non-central multivariate beta distribution and moments of traces of two matrices. *Ann. Math. Statist.* **36** 1511–1520.
- [4] KHATRI, C. G. and PILLAI, K. C. S. (1965). Further results on the non-central multivariate beta distribution and moments of traces of two matrices. Mimeograph Series No. 38, Department of Statistics, Purdue Univ.
- [5] PILLAI, K. C. S. (1954). On some distribution problems in multivariate analysis. Mimeograph Series No. 88, Institute of Statistics, Univ. of North Carolina.
- [6] PILLAI, K. C. S. (1955). Some new test criteria in multivariate analysis. *Ann. Math. Statist.* **26** 117–121.
- [7] PILLAI, K. C. S. (1956). Some results useful in multivariate analysis. *Ann. Math. Statist.* **27** 1106–1114.
- [8] PILLAI, K. C. S. (1960). *Statistical Tables for Tests of Multivariate Hypotheses*. The Statistical Center, University of the Philippines, Manila.
- [9] PILLAI, K. C. S. (1964). On the moments of elementary symmetric functions of the roots of two matrices. *Ann. Math. Statist.* **35** 1704–1712.
- [10] ROY, S. N. (1945). The individual sampling distribution of the maximum, the minimum and any intermediate of the p -statistics on the null hypotheses. *Sankhyā* **7** 133–158.