

# SOME TESTS FOR THE INTRAClass CORRELATION MODEL

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**1. Introduction.** Let  $X^{(i)}$  ( $i = 1, 2, \dots, k$ ) be independent normal random  $p$ -vectors with mean vectors  $\mu_i$  and nonsingular covariance matrices  $\Sigma_i$ . The problems with which we are concerned here are related to the comparison of dispersion matrices  $\Sigma_i$ 's, when each dispersion matrix takes the intraclass correlation form, i.e.,

$$(1) \quad \Sigma_i = \sigma_i^2[(1 - \rho_i)I + \rho_i e e'],$$

where  $e' = (1, 1, \dots, 1)$ , and  $I$  is an identity matrix of order  $p \times p$ . For this model, the problem of comparing dispersion matrices  $\Sigma_i$  reduces to that of comparing  $\sigma_i$ 's and  $\rho_i$ 's of  $k$  populations. For some related results on this model, refer to Wilks [8], Geisser [1], Votaw [7], and Selliah [6].

**2. Problems.** With the help of Roy's union-intersection principle [4], we propose in this paper test procedures for the following problems:

(i) To test the hypothesis  $H: \rho_1 = \dots = \rho_k; \sigma_1 = \dots = \sigma_k$ , against the alternative  $A: \rho_i \neq \rho_j; \sigma_i \neq \sigma_j, i, j = 1, 2, \dots, k, i \neq j$ .

(ii) To test the hypothesis  $H: \rho_i = 0, i = 1, 2, \dots, k$ , against the alternative  $A: \rho_i \neq 0, i = 1, 2, \dots, k$ .

(iii) To test the hypothesis  $H: \rho_1 = \dots = \rho_k$ , against the alternative  $A: \rho_i \neq \rho_j, i, j = 1, 2, \dots, k, i \neq j$ .

**3. Reduction to canonical form.** Let  $S$ , a  $p \times p$  matrix, have the Wishart distribution with mean  $n\Sigma$  and degrees of freedom  $n, n = N - 1$ , and let  $X$  be a normal random  $p$ -vector with mean vector  $\mu$  and covariance matrix  $\Sigma$ . In addition, let  $S$  be independent of  $X$ . For the intraclass correlation model  $\Sigma$ , there exists an orthogonal matrix  $\Gamma$  with first row  $e'/p^{1/2}$  such that  $\Gamma\Sigma\Gamma' = \text{diag}(\alpha, \beta, \dots, \beta)$ , where

$$(2) \quad \alpha = \sigma^2[1 + (p - 1)\rho], \quad \beta = \sigma^2(1 - \rho).$$

Let  $W = \Gamma S \Gamma'$ . Then the pdf of  $W$  is

$$(3) \quad p(W) = \text{Const. } \alpha^{-n/2} \beta^{-m/2} |W|^{(n-p-1)/2} \exp \left\{ -\frac{1}{2} w_{11}/\alpha - \frac{1}{2} \sum_{r=2}^p w_{rr}/\beta \right\},$$

where  $m = n(p - 1)$ . Let  $Z = \Gamma X, \eta = \Gamma\mu$ . Then  $z_1$  is  $N(\eta_1, \alpha)$ , and  $z_r$  is  $N(\eta_r, \beta)$ ,  $r = 2, \dots, p$ . The  $z_r$ 's are independently distributed for all  $r = 1, 2, \dots, p$ .

Following Olkin and Pratt [3], we can obtain statistics  $u$  and  $v$  sufficient for  $\alpha$  and  $\beta$  and distributed independently as  $\alpha\chi_a^2$  and  $\beta\chi_b^2$  where the degrees of freedom  $a$  and  $b$  depend on our knowledge of  $\mu$ . For  $\mu$  completely unknown,

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